



UNIVERSITÀ DI PARMA

ARCHIVIO DELLA RICERCA

University of Parma Research Repository

Goodness-of-fit testing for the Newcomb-Benford law with application to the detection of customs fraud

This is the peer reviewed version of the following article:

Original

Goodness-of-fit testing for the Newcomb-Benford law with application to the detection of customs fraud / Barabesi, Lucio; Cerasa, Andrea; Cerioli, Andrea; Perrotta, Domenico. - In: JOURNAL OF BUSINESS & ECONOMIC STATISTICS. - ISSN 0735-0015. - 36:2(2018), pp. 346-358. [10.1080/07350015.2016.1172014]

Availability:

This version is available at: 11381/2820297 since: 2021-11-09T15:46:42Z

Publisher:

American Statistical Association

Published

DOI:10.1080/07350015.2016.1172014

Terms of use:

Anyone can freely access the full text of works made available as "Open Access". Works made available

Publisher copyright

note finali coverpage

(Article begins on next page)

Supplementary Material for the paper:

**Goodness-of-fit testing for the
Newcomb-Benford law with application to
the detection of customs fraud**

Lucio Barabesi

University of Siena, Italy

Andrea Cerasa

European Commission, Joint Research Centre, Ispra, Italy

Andrea Cerioli

University of Parma, Italy

Domenico Perrotta

European Commission, Joint Research Centre, Ispra, Italy

March 22, 2016

Experimental results for the MAD statistic

We provide experimental results for the MAD conformity measure that complement those for the χ^2 goodness-of fit statistic given in §4. Specifically, in Table 1 we give Monte carlo estimates of the marginal and conditional quantiles of the exact distribution of the MAD statistic. These quantiles are defined in our approach as

$$q_{T_{\{l_1, \dots, l_m\}}}^*(\alpha) \tag{1}$$

and

$$q_{T_{\{l_1, \dots, l_m\}} | \Psi_{\{1, \dots, k\}}}^*(\alpha), \tag{2}$$

respectively. We take $\alpha = \beta = 0.99$.

We also report Monte Carlo estimates of the proportion of false rejections (Table 2), and comparison with the same sequential multiple-testing procedures considered in the manuscript (Table 3), using the MAD conformity measure as our test statistic.

Table 1: Monte Carlo estimates $q_{T_{\{l_1, \dots, l_m\}}}^*(0.99)$ (Marg) and $q_{T_{\{l_1, \dots, l_m\}}|\Psi_{\{1, \dots, k\}}}(0.99)$ (Cond) of the quantiles of the MAD statistic of conformity under the NB model, for $k = 1, 2, 3, 4$ and all possible index selections l_1, \dots, l_m . $B = 1,000,000$ independent replications for each sample size.

k	$\{l_1, \dots, l_m\}$	$n = 20$		$n = 100$		$n = 500$	
		Marg	Cond	Marg	Cond	Marg	Cond
2	{1}	1.7853; 2.1988		4.0485; 4.7811		9.0506; 10.7876	
	{2}	1.7055; 1.9443		3.9035; 4.6473		8.7577; 10.4095	
	{1, 2}	0.3588; –		0.9520; –		2.1303; –	
3	{1}	1.7853; 2.1988		4.0485; 4.8546		9.0506; 9.9028	
	{2}	1.7055; 1.7870		3.9035; 4.0755		8.7577; 9.2008	
	{3}	1.6141; 1.7937		3.8666; 4.0275		8.7597; 9.2928	
	{1, 2}	0.3588; 0.3772		0.9520; 1.0130		2.1303; 2.2450	
	{1, 3}	0.3591; 0.3771		0.9514; 1.0050		2.1301; 2.2333	
	{2, 3}	0.3395; 0.3419		0.8994; 0.9303		2.0889; 2.2105	
	{1, 2, 3}	0.0434; –		0.1933; –		0.6014; –	
4	{1}	1.7853; 2.1959		4.0485; 4.9374		9.0506; 10.7714	
	{2}	1.7055; 1.7510		3.9035; 4.0428		8.7577; 8.9125	
	{3}	1.6141; 1.6160		3.8666; 3.8486		8.7597; 8.9166	
	{4}	1.6015; 1.6013		3.8070; 3.9977		8.7883; 8.9844	
	{1, 2}	0.3588; 0.3771		0.9520; 1.0071		2.1303; 2.2341	
	{1, 3}	0.3591; 0.3769		0.9514; 0.9946		2.1301; 2.2063	
	{1, 4}	0.3592; 0.3766		0.9507; 0.9998		2.1305; 2.2175	
	{2, 3}	0.3395; 0.3406		0.8994; 0.9049		2.0889; 2.1076	
	{2, 4}	0.3394; 0.3401		0.8995; 0.9085		2.0888; 2.1204	
	{3, 4}	0.3400; 0.3400		0.8824; 0.8830		2.0752; 2.0906	
	{1, 2, 3}	0.0434; 0.0436		0.1933; 0.1959		0.6014; 0.6120	
	{1, 2, 4}	0.0434; 0.0436		0.1933; 0.1959		0.6015; 0.6112	
	{1, 3, 4}	0.0434; 0.0436		0.1936; 0.1961		0.6046; 0.6150	
	{2, 3, 4}	0.0393; 0.0393		0.1819; 0.1821		0.6213; 0.6248	
{1, 2, 3, 4}	0.0044; –		0.0219; –		0.1028; –		

Table 2: Monte Carlo estimates of the proportion of false rejections in testing $H_0^{\{l_1, \dots, l_m\}}$ through the MAD conformity measure when $q_{T_{\{l_1, \dots, l_m\}}}^*(0.99)$ is used instead of $q_{T_{\{l_1, \dots, l_m\}}|\Psi_{\{1, \dots, k\}}(0.99)}^*(0.99)$, for $k = 1, 2, 3, 4$ and all possible index selections l_1, \dots, l_m ($m = 1, \dots, k - 1$). For each sample size, $B = 1,000,000$ independent replications are run under the NB model and only those in which $H_0^{\{1, \dots, k\}}$ is rejected at level 0.01 are retained.

k	$\{l_1, \dots, l_m\}$	$n = 20$	$n = 100$	$n = 500$
2	{1}	0.150	0.066	0.066
	{2}	0.056	0.084	0.084
3	{1}	0.187	0.105	0.028
	{2}	0.018	0.019	0.022
	{3}	0.014	0.016	0.023
	{1, 2}	0.356	0.085	0.051
	{1, 3}	0.316	0.080	0.050
	{2, 3}	0.023	0.039	0.068
4	{1}	0.193	0.154	0.105
	{2}	0.015	0.016	0.014
	{3}	0.011	0.009	0.013
	{4}	0.009	0.012	0.013
	{1, 2}	0.333	0.072	0.042
	{1, 3}	0.288	0.061	0.036
	{1, 4}	0.297	0.058	0.041
	{2, 3}	0.015	0.013	0.013
	{2, 4}	0.014	0.014	0.019
	{3, 4}	0.011	0.012	0.014
	{1, 2, 3}	0.791	0.413	0.090
	{1, 2, 4}	0.791	0.425	0.090
	{1, 3, 4}	0.664	0.370	0.082
{2, 3, 4}	0.022	0.028	0.029	

Table 3: Monte Carlo estimates of the empirical sizes of our test using the conditional quantile $q_{T_{\{l_1, \dots, l_m\}}^* | \Psi_{\{1, \dots, k\}}(0.99)}$ (Cond), of the serial gatekeeping (Gate) and the Bonferroni-adjusted (Bonf) tests of $T_{\{1\}}$ and $T_{\{2\}}$. The MAD conformity measure is adopted and the nominal test size is 0.01. The empirical sizes are computed by conditioning on rejection of $H_0^{\{1,2\}}$, based on the estimated quantile $q_{T_{\{1,2\}}^*}^*(0.99)$. $B = 1,000,000$ independent replications are taken for each value of n .

	$n = 20$			$n = 100$			$n = 500$		
	Cond	Gate	Bonf	Cond	Gate	Bonf	Cond	Gate	Bonf
$H_0^{\{1\}}$	0.010	0.093	0.094	0.011	0.041	0.042	0.010	0.042	0.042
$H_0^{\{2\}}$	0.010	0.035	0.036	0.009	0.054	0.055	0.010	0.051	0.052